Quantifying Systemic Risk in the Presence of Unlisted Banks

Application to the European Financial Sector

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1. Motivation

- Macroprudential policy widely acknowledged to be of prime importance but implementation tends to be *ad hoc*
- A large gap between academic and policy approaches to systemic risk
 - Academic focus on implying (tail) dependencies from asset prices (Acharya eA, 2017; Adrian/Brunnermeier, 2016; ...)
 - Regulators focus on balance sheet/transaction data and regulatory scores (O-SII, G-SII scoring)

- Key challenge: many European banks are not publicly traded on the equity market
- ... but they are traded on the Credit Default Swaps (CDS) market
- Methodology is still general and could be applied to non-market data

- Implying systemic risk from market data
 - CoVaR: Adrian & Brunnermeier, 2016; SRISK: Engle, 2018;
 - MES: Acharya et al., 2017; DIP: Huang et al., 2012;
 - Lehar, 2005; Segoviano and Goodhart, 2009; Zhou,2010; [..]
- Structured Credit Risk: Merton, 1974; Leland, 1994;
- Credit Portfolio Valuation: Vasicek, 1987; Tarashev and Zhu, 2006;

Related Literature

- Financial Stability
 - Distance-to-default: Bharath and Shumway, 2008; Jessen and Lando, 2015
 - Default feedback loops: Acharya et al., 2014
- Theoretical backing
 - Fire sales: Shleifer and Vishny, 1992;
 - Correlated assets (like in Adrian/Brunnermeier (2016), Acharya e.a.(2017))
- Dimitrov/van Wijnbergen, 2023 [Macroprudential Regulation: A Risk Management Approach] develop the methodology further to calibrate banks' macroprudential capital buffers

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- CDS: insurance derivative contract (OTC) on default of an underlying
- Linked directly to the default risks of the company
 - More liquid and with fewer trading frictions than the corporate bond market
 - An edge over credit rating agencies
- Standardized T&Cs (maturities, the definition of a credit event, etc.)
- Some evidence CDS prices may lead the equity markets in price discovery *Acharya & Johnson* [2005]

- CDS prices on banks' subordinate debt: less likely to be bailed out in default
- Contract counterparty risk eliminated for centrally cleared contracts
- Less liquid than the equity market, but illiquidity often indicator of higher credit risk (Brunnermeier/Pedersen, 2009; Diamond/Rajan, 2011)
- CDS Market transparency, liquidity and resilience increased substantially since the GFC (BIS, 2018)
- Alternative sources of distress probabilities of default exist...
 - ... but how predictive are they really?

2. Model

- The regulatory space is viewed as a portfolio of loans
- Distress is defined as default on the subordinated debt of an institution
- Main idea:
 - 1. Imply default probabilities from CDS spreads
 - 2. Evaluate default correlations from CDS co-movements over time
 - 3. Evaluate the cumulative potential losses withing the system

A Model of Bank Distress

 U_i is an (unobserved) credit-worthiness variable s.t.

 $U_i \sim N(0,1)$

Default occurs if:

$$\mathbb{1}_{i} \equiv \begin{cases} 1 & \text{if } U_{i} \leq X_{i} \\ 0 & \text{otherwise} \end{cases}$$
(1)

with X_i representing a fixed default threshold (quasi-observed)

$$\implies PD_i \equiv \Phi(X_i)$$

with $\Phi(.)$ the standard cumulative normal distribution.

Default dependencies through latent factors (Gaussian Copula)

$$U_i = A_i M + \sqrt{1 - A_i A_i'} Z_i \tag{2}$$

M: Vector of stochastic common factors; Z_i : idiosyncratic factor; A_i : factor loadings, $A_iA'_i \leq 1$

• The Merton model has important implications on the implied asset correlations:

$$egin{aligned} \mathsf{a}_{ij} &\equiv \mathbb{C} ext{orr}(U_i, U_j) \ &= \mathbb{C} ext{orr}(\Delta \Phi^{-1}(-\mathsf{PD}_{i,t}), \Delta \Phi^{-1}(-\mathsf{PD}_{i,t})) \end{aligned}$$

- Determine factor exposures ρ_i to closely match these implied correlations.
- Imply time series PD_{i,t} (*) from the CDS data
- Once the exposures are fixed, the latent variables U_i can be simulated
- Default threshold is fixed by the PD-implied Distance-to-Default
 Defaults can be simulated in a multi-variate space

Quantifying Systemic Risk: (Marginal) Expected Shortfall

- The financial system can be seen as a portfolio of long loan positions
- Formally, define (correlated credit losses) as

$$L_i = \mathbb{1}_i LGD_i; \quad L_{sys} = \sum_{i=1}^N w_i L_i$$
(3)

• Define Expected Systemic Shortfall and Marginal Expected Shortfall [Acharya eA, 2017; Huang eA 2012]:

$$MES_{i} = \mathbb{E} \left(L_{i} | L_{sys} > VaR_{sys} \right)$$

$$ESS = \mathbb{E} \left(L_{sys} | L_{sys} > VaR_{sys} \right)$$
 (4)

• Percentage Contribution to ESS:

$$PC \text{ to } ESS_i = \frac{w_i MES_i}{ESS} \tag{5}$$

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3. Empirical Application

- Universe of 27 large European banks (O-SII and G-SII).
- Evaluation date: Aug, 29, 2022
- Correlation time window: 3 years
- Dataset: CDS spreads on subordinate debt; Balance sheet liability sizes

Relative Liability Size



12

Figure 1: Median Rates per Country (bps)



- Using CDS data to imply risk
 - Levels of the CDS rate speak about the market view on the credit-worthiness of the institution
 - Co-movements in CDS prices speak about the tendency of banks to be exposed to the same risk drivers
- Liability sizes speak about the Exposure at Default (EAD)

ii. Estimated Factor Loadings

		F1	F2	F3
Aus	ERST	0.93	0.02	0.04
Bel	KBCB	0.15	0.13	(0 14)
Den	DANK	0.95	0.09	0.11
Fin	NORD	0.61	<mark>(0</mark> 69)	0.20
France	SOCG	0.93	0.18	0.07
	BNP	0.96	0.20	0.06
	CRAG	0.95	0.23	0.07
	CRMU	0.51	0.09	(006)
Germany	DZ	0.86	0.01	0.10
	HESLN	0.92	(006)	0.08
	COMZ	0.95	0.16	(0.01)
	BAY	0.92	(007)	0.03
	DB	0.92	013	(008)
	LBBW	0.91	(0 <mark>.</mark> 03)	0.09
Italy	UNIC	0.92	0.11	0.05
	INTE	0.92	012	0.08
Netherlands	RABO	0.95	0 15	0.08
	ABN	0.72	0.00	(0.29)
	INGB	0.74	(007)	0.12
	VB	0.65	011	(0.21)
Spain	CAIX	0.19	(008)	<mark>(0.</mark> 49)
	SAB	0.30	(009)	<mark>(0.</mark> 64)
	SANT	0.96	0.15	(0.00)
	BBVA	0.94	0.16	(002)
Swed	SWEN	0.69	<mark>(0</mark> 62)	0.05
	SEB	0.65	<mark>(0</mark> 71)	0.02
	SWED	0.66	(0,36)	(0.28)

Joint PDs



15



Risk Contribution vs. Total CET1 Buffers



Figure 4: Expected Systemic Shortfall vs. VSTOXX



• The Student-t model

$$Ui = \sqrt{h(F)} \left(AiM + \sqrt{1 - AiAi'} Z_i \right)$$
(6)

where $h(F) = \frac{\nu}{F}$ with $F \sim \chi^2(\nu)$.

The Skewed-t model

$$Ui = \sqrt{\frac{\nu}{F}} \left(\delta G + AiM + \sqrt{1 - AiAi'} Z_i \right)$$
(7)

where $G \sim TN\left(-\sqrt{\frac{2}{\pi}},1\right)$, with $TN(\mu,\sigma)$ is a normal distribution truncated left at $-\sqrt{\frac{2}{\pi}}$.

Non-linear Factor Extensions

Figure 5: Simulated Factor Copula







Figure 9: Model Comparison PCES



Summary of findings

- *PC* to *ES* provides theoretically justified blending of risk, interdependence, and size; unlike current O-SII approaches
- Our market-based evaluation based shows large discrepancy between larger banks' capitalization and their contribution to EU-wide systemic risk
- Market-based measures of systemic risk could complement regulatory systemic rankings
- Challenges of the current O-SII methodology
 - Evidence that large banks may window dress statistics relevant for their O-SII scores
 - Significant heterogeneity across EU countries in mapping from O-SII scores to O-SII buffers
 - Systemic evaluation w.r.t local economy rather than EU-wide
- Results robust to adding non-linear factors in the Copula specification to capture asset skewness and tail-fatness

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Appendix

Appendix: Extract PDs from CDS prices

• CDS valuation, Duffie [1999]: *CDS_t* is set to equalize the expected present value of the two swap legs.

$$\underbrace{CDS_t \int_t^{t+T} e^{-r_\tau \tau} \Gamma_\tau d\tau}_{\text{PV of CDS premia}} = \underbrace{(1 - ERR_t) \int_t^{t+T} e^{-r_\tau \tau} q_\tau d\tau}_{\text{PV of protection payment}}$$
(8)

 Γ_{τ} : survival probability; r_{τ} : interest rate; *ERR*: Expected Recovery Rate; q_{τ} : hazard rate (ann. default probability, conditional on no default previously)

- Assume fixed: ERR (here only), interest rate, hazard rate
- ERR calibrated based on liabilities structure (80% on deposits/policy insurance; 40% on other)
- Set $PD_t = q_t$ for each bank $i \bowtie_{back}$

Collateral Process

• Model the value of collateral backing liabilities as:

$$d \ln C_{i,t} = \sigma_c dW_{i,t}^c \tag{9}$$

• where the collateral is defined through the factor model

$$dW_i^c = A_i M_t + \sqrt{1 - A_i A_i'} Z_{i,t}^c$$
(10)

• This generates dependent losses $(1 - RR_{i,t})$

$$RR_{i,t} = \mu_{c,i} \min(1, C_{i,t}) \tag{11}$$

RR_i: Recovery Rate

 σ_c matched to generate reasonable variance of the RRs; μ_{c,i} matched a reasonable ERR;

🕨 back

Results: Conditional



26

Relation to a Structural Credit Model

Assume the Merton firm model (under the r.n. distribution) holds

$$d\ln V_{i,t} = rdt + \sigma_i dW_{i,t} \tag{12}$$

where $V_{i,t}$ is the market value of the bank's risk-weighted assets; σ_i is their st.dev.; r is the risk-free rate; $dW_{i,t}$ is a Brownian Motion.

• Default occurs if assets fall below the face value of debt at time T

$$PD_{i,t} = \mathbb{P}(V_{i,t+T} \le D_i) \implies PD_{i,t} = \mathbb{P}\left(\underbrace{\frac{W_{i,t+T}}{\sqrt{T}}}_{U_i} \le \underbrace{-DD_{i,t}}_{X_i}\right)$$

Distance-to-Default (DD):

$$DD_{i,t} = \frac{\ln \frac{V_{i,t}}{D_i} + \left(r - \frac{\sigma_i^2}{2}\right) T}{\sigma_i \sqrt{T}}$$

➡ back

Estimate all ρ_i, ρ_j relative to a target correlation matrix

$$\min_{\rho_{i},...,\rho_{j}} \sum_{i=2}^{N} \sum_{j=1}^{N} (a_{ij} - \rho_{i}\rho_{j}')^{2}$$
(13)

with target correlations a_{ij} evaluated from co-movements in banks' PDs [Cf. Tarashev & Zhu, 2006; Andersen eA, 2003]

➡ back